

Log-Weighted Pareto Distribution And Its Statistical Properties

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ABSTRACT

The Pareto distribution is a power law probability distribution that is used to describe social scientific, geophysical, actuarial, and many other types of observable phenomena. A new weighted Pareto distribution is proposed using a logarithmic weight function. Several statistical properties of the weighted Pareto distribution are studied and derived including cumulative distribution function, location measures such as mode, median and mean, reliability measures such as reliability function, hazard and reversed hazard functions and the mean residual life, moments, shape indices such as skewness and kurtosis coefficients and order statistics. A parametric estimation is performed to obtain estimators for the distribution parameters using three different estimation methods the maximum likelihood method, the L-moments method and the method of moments. Numerical simulation is carried out to validate the robustness of the proposed distribution. The distribution is fitted to a real data set to show its importance in real life applications.

Keywords: Log-weighted Pareto distribution, weighted distributions, survival function, hazard rate function, order statistics, maximum likelihood, moment, L-moment.

1. Introduction

The theory of weighted distributions provides a collective access to the problems of model specification and data interpretation. It provides a technique for fitting models to the unknown weight functions when samples can be taken both from the original distribution and a developed distribution. Fisher [6] introduced the basic idea of the concept of weighted distributions when he studied the effect of methods of ascertainment upon estimation of frequencies. In extending the basic ideas of Fisher, Rao [16] introduced a unified theory of weighted distributions and formulated it in general terms in connection with modeling statistical data where the standard distributions were not found to be appropriate. He identified various situations that can be modeled by weighted distributions.

To introduce the concept of a weighted distribution, suppose X is a non-negative random variable with probability density function (pdf) $f(x; \theta)$, where the parameter is $\theta \in \Omega$ (Ω is the parameter space). Suppose a realization x of X under $f(x; \theta)$ enters the investigator's record with probability proportional to $w(x, \beta)$ which is a non-negative weight function with parameter β . So, the recorded x is not an observation on X , but on the random variable Z , having such pdf

$$g(z; \theta, \beta) = \frac{w(z, \beta)f(z, \theta)}{W} \quad (1)$$

where $W = E[w(x, \beta)]$ is the normalizing factor obtained to make the total probability equal to unity. The random variable Z is called the weighted version of X , and its distribution in relation to that of X is called the weighted distribution with weight function $w(x, \beta)$, Patil [15]. Different models can be obtained depending upon the choice of the weight function $w(x)$.

The concept of weighted distributions has been employed in a wide variety of applications in reliability and survival analysis, analysis of family data, meta-analysis, ecology, medicine, and forestry. The contributions of researchers to the weighted distributions vary between single and double weight distributions. Shaban and Boudrissa [19] presented a length-biased version of the Weibull distribution. Gupta and Kundu [7] developed the weighted exponential (WE) distribution as a lifetime model. Levia, et al. [11] developed a length-biased version of the Birnbaum-Saunders [BS] distribution with applications in water quality. Kersey [10] presented the size-biased inverse Weibull distribution. Shi, et al. [20] presented the theoretical properties of weighted generalized Rayleigh distribution. Hussian [9] presented the weighted inverted exponential distribution as a generalized version of the inverted exponential distribution. Seenoi, et al. [18] developed a length-biased version of the exponentiated inverted Weibull (EIW) distribution. Dey, et al. [5] and Nasiru [14] presented different forms of the Weighted Weibull Distribution. Mahmoud, et al. [12] developed the weighted Quasi-Lindley distribution and weighted Lomax distribution. Bashir and Rasul [3] introduced a new weighted Rayleigh distribution named area-biased Rayleigh distribution and they derived some of its mathematical properties. Al-kadim and Hantoosh [2] presented the double weighted exponential distribution (DWED). Ahmed and Ahmed [1] presented double weighted Rayleigh distribution (DWRD) version. Saghir and Saleem [17] presented a new version of the double weight Inverse Weibull (DWIW) distribution. Ajami and Jahanshahi [21] introduced a weighted model based on the Rayleigh distribution and they derived and studied its statistical and reliability properties. Asgharzadeh, et al. [22] introduced a generalization of Lindley distribution that provides fits for the Lindley distribution and some other distributions as special cases. Fatima and Ahmad [23] introduced weighted inverse Rayleigh distribution and investigated its different statistical properties.

2. Pareto Distribution

The Pareto distribution is named after the well-known Italian-born Swiss sociologist and economist Vilfredo Pareto (1848-1923). The Pareto distribution is a power law probability distribution that is used to describe social, scientific, geophysical, actuarial, and many other types of observable phenomena, Johnson et al. [24]. Pareto in 1896 defined his Law, which can be stated as $N = Cx^{-\alpha}$, where N represents the number of individuals in the population whose income exceeded a given level x , for some real number C and some $\alpha > 0$ [34]. Pareto distribution is commonly used in modeling heavy tailed distributions, including but not limited to income, insurance and city size populations. The Pareto distribution is often described as the basis of 80/20 rule that describes the larger compared to the smaller. A classic example is that 80% of the wealth is owned by 20% of the population. Pareto type I distribution is sometimes called the classical Pareto distribution or the European Pareto distribution. The pdf of a random variable X that follows the Pareto distribution with shape and scale parameters α , β respectively, is given by:

$$f(x; \alpha, \beta) = \frac{\alpha\beta^\alpha}{x^{\alpha+1}}, \quad x \geq \beta > 0, \alpha > 0 \quad (2)$$

During the recent decades, several new generalized Pareto distributions have been developed using modern methodologies in order to generate new families of distributions. Examples include the exponentiated Pareto distribution by Gupta et al. [25], the beta-Pareto distribution by Akinsete et al. [26] and the beta generalized Pareto distribution by Mahmoudi [27]. Alzaatreh et al. [28] defined and studied the Weibull-Pareto distribution which is a unimodal distribution and its shape can be skewed to the right or skewed to the left. Sarabia and Prieto [29] proposed Pareto positive stable distribution to study city size data. The Generalized Feller-Pareto (GFP) family was defined by Zandonatti in 2001 as mentioned in Kleiber and Kotz [30]. Odubote and Oluyede [31] defined a six-parameter class of distributions called weighted Feller-Pareto (WFP) and some related family of distributions including several other Pareto-type distributions as special cases. Hamed, et al. [32] defined and studied a generalization of the two-parameter Pareto distribution to the T-Pareto{Y} family using the T-R{Y} framework including six generalized Pareto families. Andrade and Zea [33] defined and studied a three-parameter model called the exponentiated generalized extended Pareto distribution and they provided a comprehensive mathematical treatment. In the next sections, a new distribution is proposed called Log-weighted Pareto distribution. It comes as a weighted version of the Pareto Type I distribution.

3. Log-weighted Pareto Distribution

Suppose X is a continuous random variable following the Pareto distribution with shape and scale parameters α and β respectively. Its pdf is defined as given in equation (2), as follows:

$$f(x; \alpha, \beta) = \frac{\alpha\beta^\alpha}{x^{\alpha+1}}, \quad x \geq \beta > 0, \quad \alpha > 0$$

Using Pareto as a base distribution, a new weighted distribution is introduced with $w(z) = \text{Log}(z)$ as a weight function. A random variable Z is said to have a Log-weighted Pareto (LWP) distribution, if its pdf and cdf are, respectively, given by:

$$f(z; \alpha, \beta) = \frac{\alpha^2 \beta^\alpha \log z}{1 + \alpha \log \beta z^{\alpha+1}}, \quad z \geq \beta > 1, \quad \alpha > 0 \quad (3)$$

$$F(z; \alpha, \beta) = 1 - \left(\frac{\beta}{z}\right)^\alpha \left(\frac{1 + \alpha \log z}{1 + \alpha \log \beta}\right), \quad z \geq \beta > 1, \quad \alpha > 0 \quad (4)$$

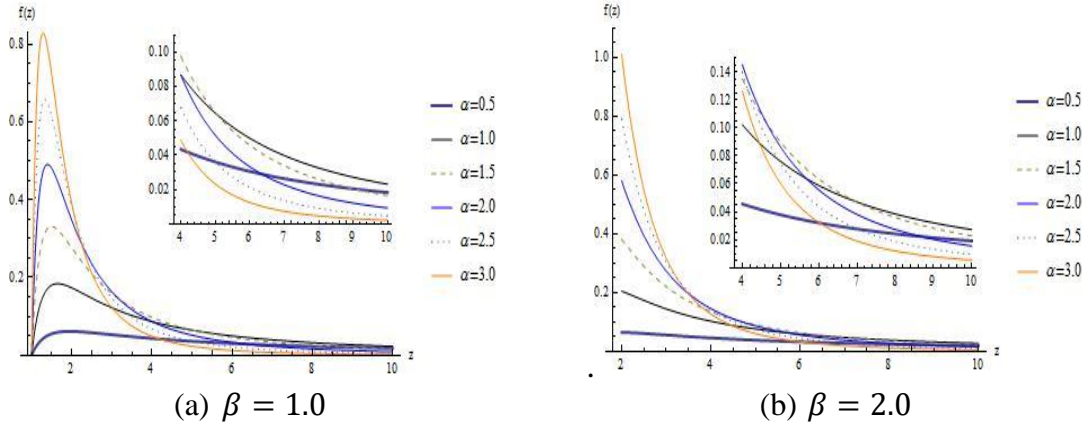


Figure (1): The LWP pdf at $\beta = 1.0$ and $\beta = 2.0$ and selected values of α

It can be easily noticed from Figure. (1a) and (1b) that at a certain value of scale parameter, β , the peak of the density function becomes higher as the value of α increases. In other words, the modal value is in inverse relation with the value of α holding the value of β constant. Moreover, the greater the value of scale parameter β , the density function becomes strictly decreasing function.

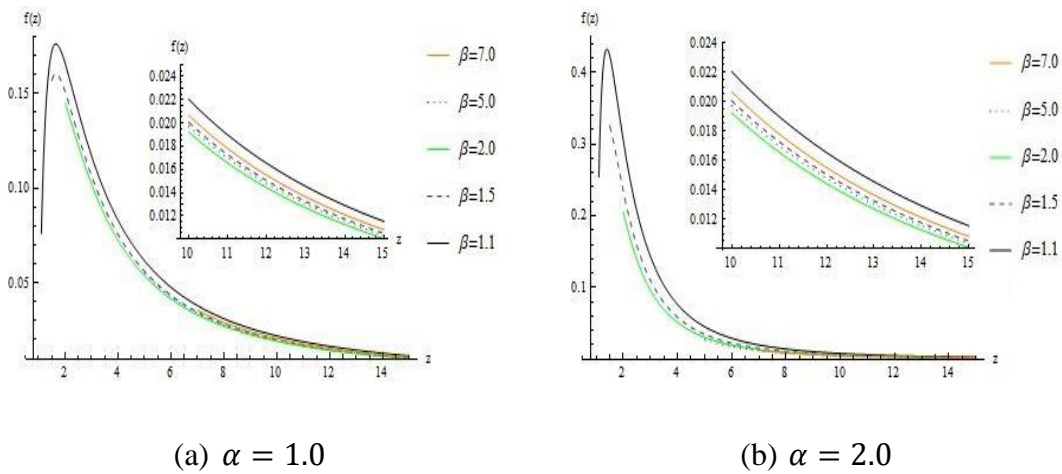
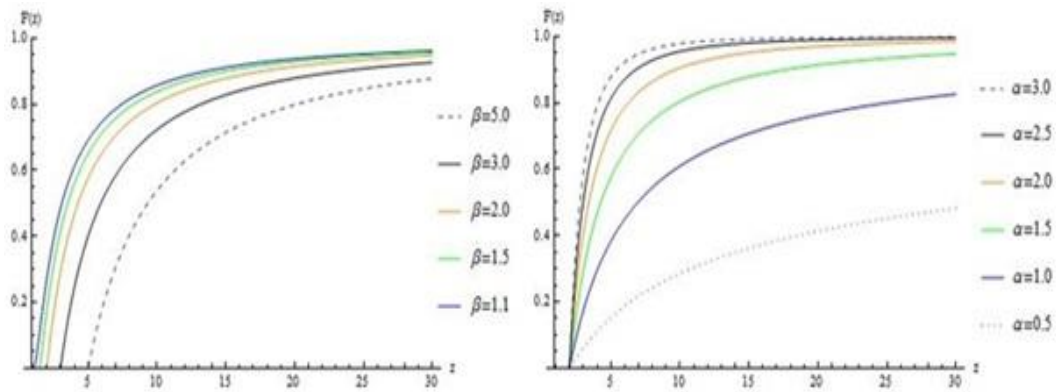


Figure (2): The LWP pdf at $\alpha = 1.0$ and $\alpha = 2.0$ and selected values of β

Similarly, it can be easily observed from Figure. (2a) and (2b) that at a certain value of the shape parameter, α , as the value of β increases, the shape of the density function is transformed from a unimodal shape to a reversed J-shape.

(a) $\alpha = 1.5$ (b) at $\beta = 2.0$ Figure (3): The LWP cdf at selected values of β and α

Also, it is clear from Figure (3a) that the smaller the value of β is, the quicker the cumulative distribution function approaches its upper bound. On the contrary to that, one can see from Figure (3b) that the smaller the value of α is, the slower the cumulative distribution function approaches its upper bound.

3.1 The Mode and Quantiles

The Log-weighted Pareto is a unimodal distribution and its modal value is achieved at $z = e^{\frac{1}{\alpha+1}}$. On the other hand, no closed form exists for the quantiles, but it can be obtained numerically by using the following relation:

$$p = F(q_p) = 1 - \left(\frac{\beta}{q_p}\right)^\alpha \left(\frac{1 + \alpha \log q_p}{1 + \alpha \log \beta}\right) \quad (5)$$

3.2 Reliability measures

The reliability function $R(\cdot)$, the hazard rate function $h(\cdot)$, the reversed hazard rate $\lambda(\cdot)$ and the mean residual life $m(\cdot)$ of the LWP distribution are given as follows:

The reliability function:

The reliability function for the Log-weighted Pareto distribution is given by:

$$R(z) = \left(\frac{\beta}{z}\right)^\alpha \left(\frac{1 + \alpha \log z}{1 + \alpha \log \beta}\right), \quad z \geq \beta > 1, \quad \alpha > 0 \quad (6)$$

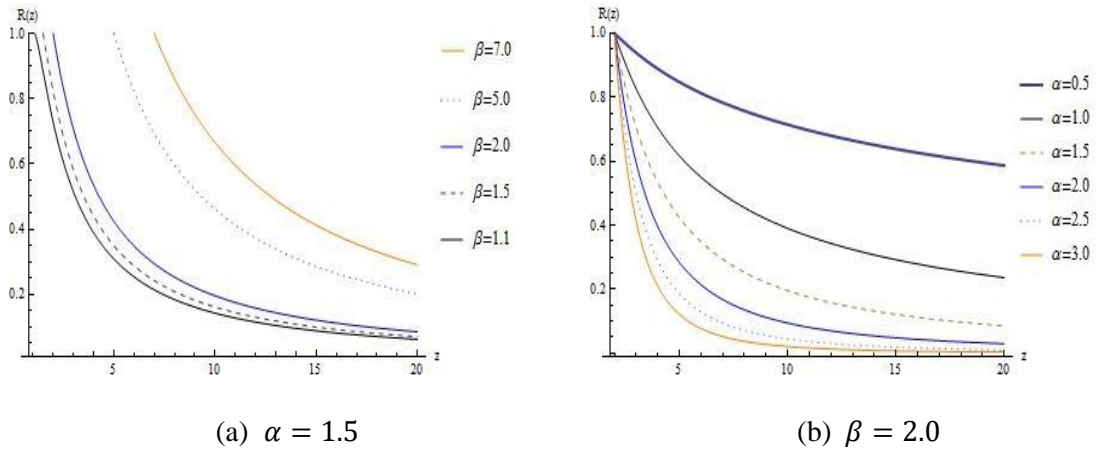


Figure (4): The LWP Reliability function at selected values of β and α

It is obvious from figure (4) that the reliability function is decreasing in z . Moreover, figure (4a) shows that at a certain value of the shape parameter α , as the scale parameter β increases, the curve of the function moves up. To the contrary to that, from figure(4b), if one holds the scale parameter β at a certain value, the curve moves down when the value of the shape parameter α increases. i.e. the reliability function has a direct relation with the scale parameter β and it has an inverse relation with the shape parameter α

The hazard function:

The hazard function for the Log-weighted Pareto distribution takes the following formula:

$$h(z) = \frac{\alpha^2 \log z}{z(1 + \alpha \log z)}, \quad z > \beta > 1, \quad \alpha > 0 \tag{7}$$

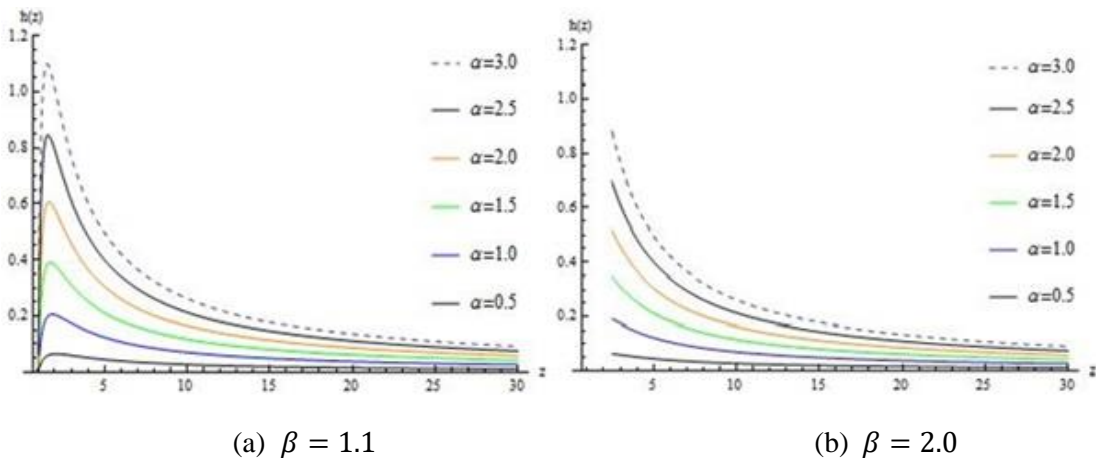


Figure (5): The hazard rate function at selected values of α

From figure (5a), one can see that, when the value of scale parameter $\beta \rightarrow 1^+$, the hazard rate increases first then it decreases, but in figure (5b), when the value of scale parameter β increases, the hazard rate is monotonically decreasing. This behavior of the hazard function can be interpreted using theorem 1 introduced by Chechile[4].The behavior of the hazard function is distributed in two ways depending on the value of the scale parameter, β . If the value of β is less than $\exp\left[\frac{\sqrt{1+4\alpha}-1}{2\alpha}\right]$, then the hazard function is a unimodal function and it reaches its peak at $z = \exp\left[\frac{\sqrt{1+4\alpha}-1}{2\alpha}\right]$, and it approaches zero as $z \rightarrow \beta^+$ or $z \rightarrow \infty$. On the other hand, if the value of β is greater than $\exp\left[\frac{\sqrt{1+4\alpha}-1}{2\alpha}\right]$, then the hazard function is

monotonically decreasing and it approaches zero as $z \rightarrow \infty$. Moreover, It is clear that the hazard function of the Log-weighted Pareto is increasing in the shape parameter, α .

The reversed hazard rate:

The reversed hazard rate is defined as the ratio of the density function to the distribution function of a random variable. The reversed hazard rate is a useful tool in the area of maintenance management. The Log-weighted Pareto distribution has the following reversed hazard function:

$$\lambda(z) = \frac{\alpha^2 \beta^\alpha z^{-1} \log z}{z^\alpha (1 + \alpha \log \beta) - \beta^\alpha (1 + \alpha \log z)}, \quad z \geq \beta > 1, \quad \alpha > 0 \quad (8)$$

It is clear from figure (6) that the reversed hazard rate is decreasing in Z . Also, one can observe from figure (6a) that, when the value of the scale parameter β increases, the curve of the reversed hazard rate function shifts to the right direction. Moreover, from figure (6b), it is obvious that as the value of the shape parameter α increases, the curve of the reversed hazard rate function moves down.

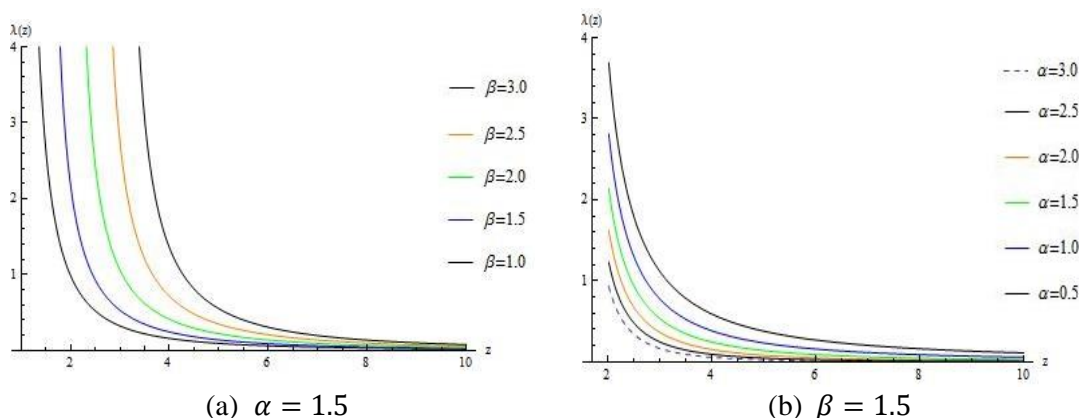


Figure (6): The reversed hazard rate at selected values of β and α

The mean residual life:

The mean residual life function measures the average residual life of a component when it has completed t units of time. It is defined by $m(t) = E(T - t | T > t)$. The mean residual life for a variable follows the Log-weighted Pareto distribution has the following form:

$$m(z) = \frac{\alpha^2 z (1 + (\alpha - 1) \log z)}{(\alpha - 1)^2 (1 + \alpha \log z)} - z, \quad z \geq \beta > 1, \quad \alpha > 1 \quad (9)$$

$$m(z) = \frac{\alpha^2 z (1 + (\alpha - 1) \log z) - z (\alpha - 1)^2 (1 + \alpha \log z)}{(\alpha - 1)^2 (1 + \alpha \log z)}, \quad z \geq \beta > 1, \quad \alpha > 1 \quad (10)$$

3.3 The moments

The r th non-central moment for a random variable Z that follows the Log-weighted Pareto is given by:

$$E(Z^r) = \frac{\alpha^2 \beta^r [1 + (\alpha - r) \log \beta]}{(\alpha - r)^2 [1 + \alpha \log \beta]}, \quad r < \alpha \quad (11)$$

Using the previous formula, the expectation and variance, coefficient of variation and coefficient of skewness can be obtained as below:

The expectation

$$E(Z) = \frac{\alpha^2 \beta [1 + (\alpha - 1) \log \beta]}{(\alpha - 1)^2 [1 + \alpha \log \beta]}, \quad \alpha > 1 \quad (12)$$

The variance

$$V(Z) = \frac{\alpha^2 \beta^2 \{1+2\alpha(\alpha-2)+(\alpha-1) \log \beta [2+4\alpha(\alpha-2)+(\alpha-2)(\alpha-1)\alpha \log \beta]\}}{(\alpha-2)^2(\alpha-1)^4(1+\alpha \log \beta)^2}, \alpha > 2 \tag{13}$$

The coefficient of variation

$$CV = \frac{\sqrt{\{1+2\alpha(\alpha-2)+(\alpha-1) \log \beta [2+4\alpha(\alpha-2)+(\alpha-2)(\alpha-1)\alpha \log \beta]\}}}{\alpha(\alpha-2)(1+(\alpha-1) \log \beta)}, \alpha > 2 \tag{14}$$

Table 1 Mean and variance of LWPDP for some values of α and β

β	1.5		2		2.5		5		10	
α	Mean	Variance	Mean	Variance	Mean	Variance	Mean	Variance	Mean	Variance
2.5	3.3277	22.524	4.1465	32.080	5.0108	44.128	9.4392	135.43	18.3112	460.71
3	2.7576	5.2367	3.4871	7.6338	4.2502	10.689	8.1434	34.421	15.9484	121.52
4	2.2543	1.1346	2.9023	1.6973	3.5715	2.4238	6.9654	8.2051	13.7686	30.013
7	1.8260	0.1440	2.3998	0.2238	2.9822	0.3284	5.9126	1.1826	11.7803	4.4970

From the previous table, it is obvious that the values of mean and variance are decreasing in α , but they are increasing in β . If the values of the coefficient of variation are calculated for these combinations of values of the parameters α and β , one can observe that the value of the coefficient of variation is decreasing in both α and β .

Index of Skewness:

The coefficient of skewness is given by:

$$\sqrt{\beta_1} = \frac{\mu_3}{\sqrt{\mu_2^3}}, \quad \alpha > 3, \quad \beta > 1 \tag{15}$$

where μ_2 and μ_3 are given by:

$$\mu_2 = \frac{\alpha^2 \beta^2 \{1+2\alpha(\alpha-2)+(\alpha-1) \log \beta [2+4\alpha(\alpha-2)+(\alpha-2)(\alpha-1)\alpha \log \beta]\}}{(\alpha-2)^2(\alpha-1)^4(1+\alpha \log \beta)^2} \tag{16}$$

$$\mu_3 = \frac{2\alpha^2 \beta^3 (2 + A_1 + [(\alpha - 1) \log \beta] \{6 + 3A_1 + [\alpha(\alpha - 1) \log \beta] A_2\})}{(\alpha - 3)^2 (\alpha - 2)^2 (\alpha - 1)^6 (1 + \alpha \log \beta)^3} \tag{17}$$

where A_1 and A_2 are as follow:

$$A_1 = \alpha(\alpha(\alpha(2\alpha - 5) - 10) + 29) - 14 \tag{18}$$

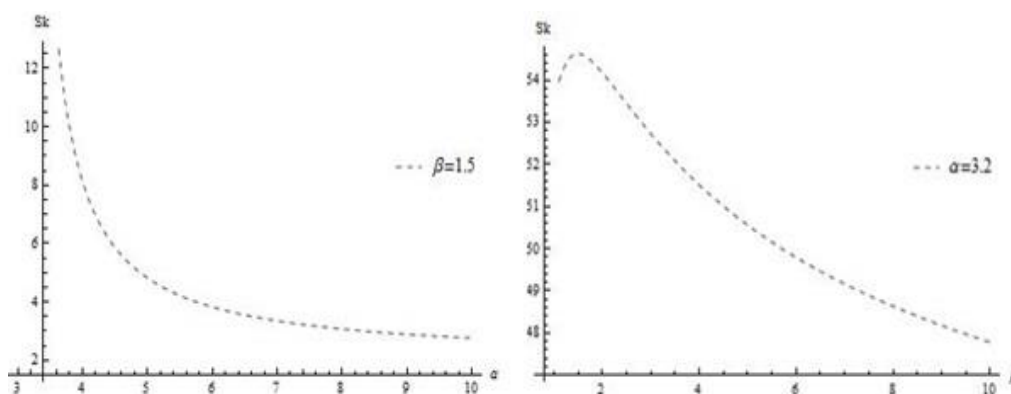
$$A_2 = 3(2\alpha^4 - 7\alpha^3 + 13\alpha - 4) + (\alpha - 3)(\alpha - 2)(\alpha - 1)\alpha(\alpha + 1) \log \beta \tag{19}$$

The value of the coefficient of skewness can be calculated for several values of the parameters α and β which presented in the following table. It is obvious from the values in the table that all the values are positive. Moreover, this coefficient is decreasing in α . For $\alpha \leq 3$, another formula may be used instead of the moment formula.

Table2: Skewness coefficient at several combinations of values of α and β

α	β	1.0	1.2	1.5	2.0	2.5	3.0	5.0
3.2		53.7293	54.2818	54.6231	54.1713	53.4397	52.7244	50.5569
3.5		15.7919	16.0756	16.3608	16.4416	16.3796	16.2881	15.9557
3.8		9.5001	9.7335	9.9766	10.0911	10.0964	10.0734	9.9545
4.0		7.6755	7.8967	8.1242	8.2405	8.2590	8.2508	8.1808
4.2		6.5312	6.7470	6.9636	7.0783	7.1028	7.1022	7.0585
4.5		5.4413	5.6555	5.8612	5.9716	5.9997	6.0047	5.9820

4.8	4.7426	4.9595	5.1577	5.2634	5.2923	5.2998	5.2882
5.0	4.4009	4.6208	4.8151	4.9177	4.9466	4.9550	4.9479



(a) $\beta = 1.5$

(b) $\alpha = 3.2$

Figure (7): The coefficient of skewness versus the values of α and β

From figure (7a), one can observe that the value of the coefficient of skewness value decreases sharply for the value of α , holding the value of β constant. On the other hand, the skewness coefficient value increases and then decreases for the value of β , keeping the value of α at the same level.

Index of Kurtosis:

The coefficient of kurtosis of the Log-weighted Pareto distribution takes the formula:

$$\beta_2 = \frac{\mu_4}{\mu_2^2} \tag{20}$$

where

$$\mu_4 = \frac{A}{B}, \quad \alpha > 4 \tag{21}$$

$$A = 3\alpha^2\beta^4(A_1 + 4A_2 + 2A_3 + 4A_4 + A_5) \tag{22}$$

$$A_1 = 12 - 116\alpha + 423\alpha^2 - 694\alpha^3 + 503\alpha^4 - 288\alpha^5 + 180\alpha^6 - 64\alpha^7 + 8\alpha^8 \tag{23}$$

$$A_2 = \log \beta (128\alpha - 539\alpha^2 + 1117\alpha^3 - 1197\alpha^4 + 791\alpha^5 - 468\alpha^6 + 244\alpha^7 - 72\alpha^8 + 8\alpha^9 - 12) \tag{24}$$

$$A_3 = \alpha(\alpha - 1)^2 \log^2 \beta (460\alpha - 934\alpha^2 + 649\alpha^3 - 500\alpha^4 + 439\alpha^5 - 172\alpha^6 + 22\alpha^7 - 72) \tag{25}$$

$$A_4 = \alpha^2(\alpha - 1)^3 \log^3 \beta (-124\alpha + 89\alpha^2 - 118\alpha^3 + 123\alpha^4 - 48\alpha^5 + 6\alpha^6 + 36) \tag{26}$$

$$A_5 = \alpha^3(\alpha - 1)^4 \log^4 \beta (28\alpha - 64\alpha^2 + 71\alpha^3 - 26\alpha^4 + 3\alpha^5 - 48) \tag{27}$$

$$B = (\alpha - 4)^2(\alpha - 3)^2(\alpha - 2)^2(\alpha - 1)^8(1 + \alpha \log \beta)^4 \tag{28}$$

The value of the coefficient of kurtosis can be calculated for several values of the two distribution parameters α and β which is presented in the following table. It is obvious from the values in the table that the coefficient is decreasing sharply in α for $\alpha > 4$. For the interval $\alpha \leq 4$, another formula may be used instead of the moment formula.

Table3: Kurtosis coefficient at several combinations of values of α and β

α	β	1.0	1.2	1.5	2.0	2.5	3.0	5.0
4.3		444.326	467.836	490.022	497.141	494.035	488.829	469.658
4.5		202.803	215.169	227.256	232.556	232.559	231.268	225.237
4.8		104.021	111.461	118.720	122.375	122.995	122.802	120.895
5.0		77.300	83.332	89.131	92.151	92.801	92.802	91.755
5.5		46.797	51.162	55.143	57.258	57.820	57.949	57.656
7.0		22.641	25.675	27.978	29.122	29.456	29.573	29.605

3.4 Order Statistics

Let Z be a random variable following the Log-weighted Pareto distribution with shape and scale parameters α and β respectively. The density function of the k^{th} order statistic, $Z_{(k)}$, from a sample of size n has the following form:

$$f_{k:n}(z) = \frac{n!}{(k-1)!(n-k)!} \frac{\alpha^2 \beta^{(n-k+1)\alpha}}{1 + \alpha \log \beta} \left(1 - \left(\frac{\beta}{z}\right)^\alpha \left(\frac{1 + \alpha \log z}{1 + \alpha \log \beta}\right)\right)^{k-1} \left(\frac{1 + \alpha \log z}{1 + \alpha \log \beta}\right)^{n-k} \frac{\log z}{z^{(n-k+1)\alpha+1}} \quad (29)$$

The minimum and the maximum order statistics probability density functions are as follow:

$$f_{1:n}(z) = \frac{n\alpha^2 \beta^{n\alpha}}{(1 + \alpha \log \beta)^n} \frac{\log z}{z^{n\alpha+1}} (1 + \alpha \log z)^{n-1}, \quad z \geq \beta > 1, \alpha > 0 \quad (30)$$

$$f_{n:n}(z) = n \frac{\alpha^2 \beta^\alpha}{1 + \alpha \log \beta} \frac{\log z}{z^{\alpha+1}} \left(1 - \left(\frac{\beta}{z}\right)^\alpha \left(\frac{1 + \alpha \log z}{1 + \alpha \log \beta}\right)\right)^{n-1}, \quad z \geq \beta > 1, \alpha > 0 \quad (31)$$

A special case

If $n=2$, the pdf, cdf and the expectation of the first and the second order statistics are:

$$f_{1:2}(z) = \frac{2\alpha^2 \beta^{2\alpha}}{(1 + \alpha \log \beta)^2} \frac{\log z}{z^{2\alpha+1}} (1 + \alpha \log z), \quad z \geq \beta > 1, \quad \alpha > 0 \quad (32)$$

$$F_{1:2}(z) = 1 - \frac{\beta^{2\alpha}}{(1 + \alpha \log \beta)^2} \frac{(1 + \alpha \log z)^2}{z^{2\alpha}}, \quad z \geq \beta > 1, \quad \alpha > 0 \quad (33)$$

$$E(Z_{1:2}) = \frac{2\alpha^2 \beta [\{1 - 6\alpha + 8\alpha^2 + (1 - 2\alpha)^2 \alpha \log \beta\} \log \beta + 4\alpha - 1]}{(2\alpha - 1)^3 (1 + \alpha \log \beta)^2}, \quad \alpha > \frac{1}{2} \quad (34)$$

$$f_{2:2}(z) = \frac{2\alpha^2 \beta^\alpha}{1 + \alpha \log \beta} \frac{\log z}{z^{\alpha+1}} \left(1 - \left(\frac{\beta}{z}\right)^\alpha \left(\frac{1 + \alpha \log z}{1 + \alpha \log \beta}\right)\right), \quad z \geq \beta > 1, \quad \alpha > 0 \quad (35)$$

$$F_{2:2}(z) = \left[1 + \frac{\left(\frac{\beta}{z}\right)^\alpha (1 + \alpha \log z) \left[\left(\frac{\beta}{z}\right)^\alpha (1 + \alpha \log z) - 2(1 + \alpha \log \beta)\right]}{(1 + \alpha \log \beta)^2}\right], \quad z \geq \beta \quad (36)$$

$$E(Z_{2:2}) = \frac{2\alpha^4 \beta [\log \beta \{(\alpha - 1)(1 - 2\alpha)^2 \log \beta + 2\alpha(4\alpha - 5) + 3\} + 4\alpha - 3]}{(\alpha - 1)^2 (2\alpha - 1)^3 (1 + \alpha \log \beta)^2}, \quad \alpha > 1 \quad (37)$$

4. Parameter Estimation

The Log-weighted Pareto distribution has two parameters. In order to obtain estimators for these parameters, three different methods are used here, namely, moments method, L-moments method and the maximum likelihood method.

4.1 The method of moments

The method of moments is a simple technique based on the idea that the sample moments are “natural” estimators of population moments. If Z_1, \dots, Z_n are assumed to be independent and identically distributed then the estimators of the distribution parameters $\theta_1, \dots, \theta_p$ are obtained by solving the set of p equations:

$$\mu'_k = m'_k \quad , \quad k = 1, 2, \dots, p \quad (38)$$

where μ'_k is the k^{th} population moment and m'_k is the k^{th} sample moment.

Applying the method of moments on the Log-weighted Pareto distribution, the following equations are obtained:

$$\frac{\sum z_i}{n} = \frac{\alpha^2 \beta [1 + (\alpha - 1) \log \beta]}{(\alpha - 1)^2 [1 + \alpha \log \beta]} \quad , \quad \alpha > 1 \quad (39)$$

$$\frac{\sum z_i^2}{n} = \frac{\alpha^2 \beta^2 [1 + (\alpha - 2) \log \beta]}{(\alpha - 2)^2 [1 + \alpha \log \beta]} \quad , \quad \alpha > 2 \quad (40)$$

Solving equations (39) and (40) simultaneously, we obtain $\hat{\beta}$ and $\hat{\alpha}$.

4.2 The method of L-moments

L-moments have some theoretical advantages over conventional moments of being more robust to the presence of outliers in the data. Experience also shows that, compared with conventional moments, L-moments are less subject to bias in estimation and approximate their asymptotic normal distribution more closely in finite samples. Parameter estimates obtained from L-moments are some times more accurate in small.

The first two L-moments are that defined by:

$$l_1 = E[Z_{1:1}] \quad (41)$$

$$l_2 = \frac{1}{2} E[Z_{2:2} - Z_{1:2}] \quad (42)$$

In the Log-weighted Pareto distribution, the previous L-moments are given by:

$$l_1 = \frac{\alpha^2 \beta (1 + (\alpha - 1) \log \beta)}{(\alpha - 1)^2 (1 + \alpha \log \beta)} \quad , \quad \alpha > 1 \quad (43)$$

$$l_2 = \frac{\alpha^2 \beta ((12\alpha^3 - 18\alpha^2 + 8\alpha - 1) \log \beta + (1 - 2\alpha)^2 (\alpha - 1) \alpha \log^2 \beta + 6\alpha^2 - 6\alpha + 1)}{(\alpha - 1)^2 (2\alpha - 1)^3 (1 + \alpha \log \beta)^2} \quad , \quad \alpha > 1 \quad (44)$$

The following two equations can be solved simultaneously to obtain the estimators for the two parameters α and β .

$$l_1 = \bar{Z} \quad (45)$$

$$l_2 = \frac{1}{2} \binom{n}{2}^{-1} \sum_i \sum_j (z_{i:n} - z_{j:n}) \quad , \quad i > j \quad , \quad \alpha > 1 \quad (46)$$

4.3 The maximum likelihood method

The likelihood function is viewed as a function of the parameters α and β as follows:

$$L(\mathbf{z}; \alpha, \beta) = \prod_{i=1}^n \frac{\alpha^2 \beta^\alpha}{1 + \alpha \log \beta} \frac{\log z_i}{z_i^{\alpha+1}} = \frac{\alpha^{2n} \beta^{n\alpha}}{(1 + \alpha \log \beta)^n} \prod_{i=1}^n z_i^{-(\alpha+1)} \log z_i \quad (47)$$

The Log-likelihood function is given by:

$$\log L = \log \left[\frac{\alpha^{2n} \beta^{n\alpha}}{(1 + \alpha \log \beta)^n} \right] + \sum_{i=1}^n \log(z_i^{-(\alpha+1)} \log z_i) \quad (48)$$

$$\log L = 2n \log \alpha + n\alpha \log \beta - n \log(1 + \alpha \log \beta) - (\alpha + 1) \sum_{i=1}^n \log z_i + \sum_{i=1}^n \log(\log z_i) \quad (49)$$

It is clear that the previous function is increasing in β , so the value of β that maximizes this function is the first order statistic $Z_{(1)}$, then $\tilde{\beta} = Z_{(1)}$ is the maximum likelihood estimator of β . Moreover, using the partial derivative of the likelihood function with respect to α , one can obtain the following equation:

$$\frac{2n}{\alpha} + n \log \beta - \frac{n \log \beta}{(1 + \alpha \log \beta)} - \sum_{i=1}^n \log z_i = 0 \quad (50)$$

Substitute $\tilde{\beta} = Z_{(1)}$ in (50), we obtain the maximum likelihood estimator, $\tilde{\alpha}$ for α .

4.4 Simulation study

In this section, a study of the behavior of the estimators for the unknown parameters α and β is considered. The estimation is made when the two parameters are unknown. Three different methods were used: the maximum likelihood method, the L-moments method and method of moments. Several combinations of the values of the two parameters are assumed. The shape parameter α is assumed to take the values 0.2, 0.5, 1.0, 1.5, 2.0, 2.5, 3.0 and 5.0, where the scale parameter β is as 1.2, 1.5, 2.0, 2.5, 3.0 and 5.0. The estimated values, biases, relative biases, mean squared error (MSE) and the scaled mean squared error were computed. This is done by generating samples from the Log-weighted Pareto distribution and considering samples of sizes of 30, 50, 100, 150 and 200. The simulations are based on 10000 replications. There were some constraints on using the estimation methods according to the value of the shape parameter, α . For the interval $0 < \alpha \leq 1$, only the maximum likelihood method is applicable. For the interval $1 < \alpha \leq 2$, both the maximum likelihood and L-moments methods are applicable. For the interval $\alpha > 2$, all of the three considered estimation methods are applied. The following tables present some of the obtained results.

The maximum likelihood method is applied for estimating the two parameters in the interval $\alpha \in (0, 1]$. From tables 4, 5 and 6, it can be observed bias, relative bias, MSE and scaled MSE values of both parameters α and β decrease as the sample size increases. Moreover, the values of relative bias and scaled MSE are also decreasing in α . In addition, the values of bias and MSE for α are always less than the corresponding values for β in this interval.

The maximum likelihood method and the L-moments method are applied for estimating the two parameters in the interval, $\alpha \in (1, 2]$. From table 7 through table 11, bias, relative bias, MSE and scaled MSE values of both parameters α and β decrease as the sample size increases. Moreover, the smaller the values of α , the less the values of bias and MSE for α than the corresponding values for β . The larger the values of α , the less the values of bias and MSE for β than the corresponding values for α . In addition, it is clear that the maximum likelihood estimators for both parameters are less bias and more efficient than those of the L-moments either for small sample sizes or large sample sizes.

From table 12 through table 15, it can be easily observed that the values of bias, relative bias, MSE and scaled root MSE of both parameters α and β decrease as the sample size increases for all estimation methods. Moreover, the values of bias and MSE for β are less than the corresponding values for α . In addition, it is clear that the maximum likelihood

estimators for both parameters are less bias and more efficient than the other methods either for small sample sizes or large sample sizes.

To sum up, it is clear that the estimators obtained by the maximum likelihood method are the most efficient among the other used estimation methods. However, the L-moments method gives better estimates than the moment method which appears in the corresponding values of bias and MSE. All the estimates improved in the values of bias, relative bias, MSE and scaled root MSE as the sample size increases, but the maximum likelihood estimates still the best. Moreover, the estimates of the shape parameter α have less bias and MSE values than the estimates of the scale parameter β for small values of the shape parameter α , but the opposite becomes true for the large values of α .

5. Application to Real Data

A real data set is used to compare the fits of the Log-weighted Pareto distribution with the Pareto distribution. This data set represents the failure times of 24 Mechanical Components which is mentioned in Murthy et. al, [13]. The data set is reported in the following table:

30.94 18.51 16.62 51.56 22.85 22.38 19.08 49.56
 17.12 10.67 25.43 10.24 27.47 14.70 14.10 29.93
 27.98 36.02 19.40 14.97 22.57 12.26 18.14 18.84

In the following table, the parameters of the distributions are estimated using the maximum likelihood method. For comparison, the Kolmogorov-Smirnov statistic is considered where the lower value for this statistic indicates a good fit. The computations of this statistic are carried out using mathematica software. The results are listed in the following table

Table: Parameter estimates and K-S statistics for Failure Times of 24 Mechanical Components

Distribution	Pareto	Log-weighted Pareto
Maximum likelihood estimates	$\hat{\alpha} = 1.394$ $\hat{\beta} = 10.240$	$\hat{\alpha} = 1.678$ $\hat{\beta} = 10.240$
K-S statistics	0.24083	0.23276
P-value	0.10397	0.12593

6. Conclusion

In this paper, the Log-weighted Pareto distribution is proposed. A mathematical treatment of the proposed distribution including explicit formulas for the density and hazard functions, moments, order statistics have been provided. The estimation of the parameters has been approached by maximum likelihood method and method of moments and the L-moments method. A simulation study based on 10000 replications was applied for estimating the distribution parameters by several sample sizes.

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Appendix 1: Tables

Table 4: Biases and MSEs of $\hat{\alpha}$ and $\hat{\beta}$ using ML method at $\beta = 1.2$ and $\beta = 1.5$ and multiple α values

$\beta = 1.2$									
α	n	$\hat{\alpha}$				$\hat{\beta}$			
		Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE
0.2	30	0.06949	0.34745	0.00567	0.37650	2.99495	2.49579	26.9198	4.32369
	50	0.06350	0.31750	0.00448	0.33466	1.61560	1.34633	5.70443	1.99033
	100	0.05957	0.29785	0.00376	0.30659	0.82954	0.69128	1.26013	0.93546
	150	0.05802	0.29010	0.00350	0.29580	0.56479	0.47066	0.55729	0.62210
	200	0.05737	0.28685	0.00339	0.29112	0.43197	0.35998	0.31937	0.47094
0.5	30	0.02629	0.05258	0.00577	0.15192	0.59976	0.49980	0.68319	0.68879
	50	0.01606	0.03212	0.00306	0.11063	0.37941	0.31618	0.25322	0.41934
	100	0.00776	0.01552	0.00136	0.07376	0.20706	0.17255	0.07581	0.22945
	150	0.00538	0.01076	0.00087	0.05899	0.12130	0.10108	0.03983	0.16631
	200	0.00469	0.00938	0.00066	0.05138	0.11403	0.09503	0.02308	0.12660
1.0	30	0.05052	0.05052	0.02304	0.15179	0.20036	0.16697	0.07304	0.22522
	50	0.02842	0.02842	0.01203	0.10968	0.12603	0.10503	0.02919	0.14238
	100	0.01353	0.01353	0.00558	0.07470	0.06749	0.05624	0.00832	0.07601
	150	0.00847	0.00847	0.00361	0.06008	0.04666	0.03888	0.00405	0.05303
	200	0.00679	0.00679	0.00271	0.05206	0.03505	0.02921	0.00230	0.03997
$\beta = 1.5$									
α	N	$\hat{\alpha}$				$\hat{\beta}$			
0.2	30	0.06224	0.31120	0.00469	0.34242	2.86391	1.90927	22.4779	3.16072
	50	0.05628	0.28140	0.00362	0.30083	1.61731	1.07821	6.62394	1.71580
	100	0.05241	0.26205	0.00296	0.27203	0.76376	0.50917	1.18160	0.72468
	150	0.05143	0.25715	0.00279	0.26410	0.50830	0.33887	0.49767	0.47030
	200	0.05069	0.25345	0.00267	0.25836	0.38419	0.25613	0.28715	0.35724
0.5	30	0.02613	0.05226	0.00575	0.15166	0.53792	0.35861	0.57776	0.50674
	50	0.01638	0.03276	0.00313	0.11189	0.33119	0.22079	0.21343	0.30799
	100	0.00775	0.01550	0.00139	0.07457	0.16398	0.10932	0.05139	0.15113
	150	0.00612	0.01224	0.00091	0.06033	0.11247	0.07498	0.02467	0.10471
	200	0.00439	0.00878	0.00066	0.05138	0.08489	0.05659	0.01400	0.07888
1.0	30	0.04871	0.04871	0.02281	0.15103	0.16743	0.11162	0.05485	0.15613
	50	0.03181	0.03181	0.01274	0.11287	0.10129	0.06753	0.02028	0.09494
	100	0.01468	0.01468	0.00569	0.07543	0.04941	0.03294	0.00484	0.04638
	150	0.00899	0.00899	0.00367	0.06058	0.03398	0.02265	0.00226	0.03169
	200	0.00699	0.00699	0.00275	0.05244	0.02611	0.01741	0.00136	0.02459

Table 5: Biases and MSEs of $\hat{\alpha}$ and $\hat{\beta}$ using ML method at $\beta = 2.0$ and $\beta = 2.5$ and multiple α values

$\beta = 2.0$									
α	n	$\hat{\alpha}$				$\hat{\beta}$			
		Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE
0.2	30	0.06514	0.32570	0.00507	0.35602	2.92396	1.46198	29.6099	2.72075
	50	0.05958	0.29790	0.00399	0.31583	1.54829	0.77415	5.83030	1.20730
	100	0.05588	0.27940	0.00334	0.28896	0.72460	0.36230	1.10789	0.52628
	150	0.05462	0.27310	0.00312	0.27928	0.47500	0.23750	0.46014	0.33917
	200	0.05399	0.26995	0.00302	0.27477	0.35646	0.17823	0.25864	0.25428
0.5	30	0.02531	0.05062	0.00581	0.15245	0.55363	0.27682	0.68106	0.41263
	50	0.01509	0.03018	0.00307	0.11082	0.31571	0.15786	0.20757	0.22780
	100	0.00787	0.01574	0.00140	0.07483	0.15626	0.07813	0.04938	0.11111
	150	0.00534	0.01068	0.00091	0.06033	0.10225	0.05113	0.02100	0.07246
	200	0.00428	0.00856	0.00069	0.05254	0.07839	0.03920	0.01232	0.05550
1.0	30	0.05012	0.05012	0.02522	0.15881	0.16559	0.08280	0.05617	0.11850
	50	0.02983	0.02983	0.01350	0.11619	0.09939	0.04970	0.01992	0.07057
	100	0.01326	0.01326	0.00591	0.07688	0.04932	0.02466	0.00485	0.03482
	150	0.00951	0.00951	0.00382	0.06181	0.03247	0.01624	0.00208	0.02280
	200	0.00556	0.00556	0.00291	0.05394	0.02457	0.01229	0.00121	0.01739
$\beta = 2.5$									
α	n	$\hat{\alpha}$				$\hat{\beta}$			
		Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE
0.2	30	0.04361	0.21805	0.00267	0.25836	3.43619	1.37448	37.7775	2.45854
	50	0.03855	0.19275	0.00189	0.21737	1.81090	0.72436	8.42220	1.16084
	100	0.03476	0.17380	0.00141	0.18775	0.81352	0.32541	1.45446	0.48240
	150	0.03381	0.16905	0.00127	0.17819	0.51999	0.20800	0.57840	0.30421
	200	0.03325	0.16625	0.00120	0.17321	0.39199	0.15680	0.32089	0.22659
0.5	30	0.02564	0.05128	0.00597	0.15453	0.57422	0.22969	0.72549	0.34070
	50	0.01538	0.03076	0.00313	0.11189	0.33574	0.13430	0.23723	0.19483
	100	0.00765	0.01530	0.00145	0.07616	0.16379	0.06552	0.05527	0.09404
	150	0.00532	0.01064	0.00092	0.06066	0.10823	0.04329	0.02373	0.06162
	200	0.00443	0.00886	0.00069	0.05254	0.08023	0.03209	0.01274	0.04515
1.0	30	0.05037	0.05037	0.02525	0.15890	0.17811	0.07124	0.06706	0.10358
	50	0.03114	0.03114	0.01355	0.11640	0.10599	0.04240	0.02361	0.06146
	100	0.01402	0.01402	0.00594	0.07707	0.05262	0.02105	0.00557	0.02985
	150	0.00845	0.00845	0.00396	0.06293	0.03536	0.01414	0.00250	0.02000
	200	0.00691	0.00691	0.00295	0.05431	0.02601	0.01040	0.00136	0.01475

Table 6: Biases and MSEs of $\hat{\alpha}$ and $\hat{\beta}$ using ML method at $\beta = 3.0$ and $\beta = 5.0$ and multiple α values

$\beta = 3$									
		$\hat{\alpha}$				$\hat{\beta}$			
α	n	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE
0.2	30	0.03639	0.18195	0.00207	0.22749	3.93984	1.31328	60.1469	2.58515
	50	0.03148	0.15740	0.00139	0.18641	1.98353	0.66118	10.86760	1.09887
	100	0.02839	0.14195	0.00099	0.15732	0.87230	0.29077	1.75114	0.44110
	150	0.02729	0.13645	0.00087	0.14748	0.55955	0.18652	0.67272	0.27340
	200	0.02664	0.13320	0.00080	0.14142	0.41179	0.13726	0.34929	0.19700
0.5	30	0.02658	0.05316	0.00595	0.15427	0.60954	0.20318	0.84532	0.30647
	50	0.01549	0.03098	0.00323	0.11367	0.35288	0.11763	0.26177	0.17054
	100	0.00744	0.01488	0.00141	0.07510	0.17670	0.05890	0.06319	0.08379
	150	0.00569	0.01138	0.00095	0.06164	0.11586	0.03862	0.02734	0.05512
	200	0.00394	0.00788	0.00069	0.05254	0.08689	0.02896	0.01516	0.04104
1.0	30	0.05203	0.05203	0.02646	0.16267	0.20155	0.06718	0.08443	0.09686
	50	0.02876	0.02876	0.01378	0.11739	0.11754	0.03918	0.02835	0.05612
	100	0.01335	0.01335	0.00629	0.07931	0.05794	0.01931	0.00689	0.02767
	150	0.00879	0.00879	0.00401	0.06332	0.03822	0.01274	0.00293	0.01804
	200	0.00702	0.00702	0.00302	0.05495	0.02881	0.00960	0.00168	0.01366
$\beta = 5$									
α	n	$\hat{\alpha}$				$\hat{\beta}$			
0.2	30	0.08477	0.42385	0.00813	0.45083	3.98026	0.79605	53.1158	1.45761
	50	0.07919	0.39595	0.00678	0.41170	2.02662	0.40532	10.80270	0.65735
	100	0.07477	0.37385	0.00584	0.38210	0.93435	0.18687	1.88987	0.27495
	150	0.07312	0.36560	0.00550	0.37081	0.59808	0.11962	0.75709	0.17402
	200	0.07255	0.36275	0.00538	0.36674	0.43429	0.08686	0.38947	0.12482
0.5	30	0.03249	0.06498	0.00652	0.16149	0.83524	0.16705	1.57224	0.25078
	50	0.02092	0.04184	0.00339	0.11645	0.47226	0.09445	0.47052	0.13719
	100	0.01382	0.02764	0.00160	0.08000	0.22844	0.04569	0.10825	0.06580
	150	0.01156	0.02312	0.00106	0.06512	0.15193	0.03039	0.04723	0.04346
	200	0.01036	0.02072	0.00078	0.05586	0.11292	0.02258	0.02622	0.03239
1.0	30	0.05243	0.05243	0.02855	0.16897	0.28071	0.05614	0.16475	0.08118
	50	0.03069	0.03069	0.01506	0.12272	0.16979	0.03396	0.05854	0.04839
	100	0.01597	0.01597	0.00686	0.08283	0.08303	0.01661	0.01404	0.02370
	150	0.01017	0.01017	0.00447	0.06686	0.05431	0.01086	0.00603	0.01553
	200	0.00586	0.00586	0.00321	0.05666	0.04086	0.00817	0.00337	0.01161

Table 7: Biases and MSEs of $\hat{\alpha}$ and $\hat{\beta}$ using ML and L-M methods at $\beta = 1.2$ and multiple values of α

$\beta = 1.2$																	
Maximum likelihood										L-moments							
$\hat{\alpha}$					$\hat{\beta}$				$\hat{\alpha}$				$\hat{\beta}$				
a	n	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE
1	30	0.0	0.0	0.0	0.1	0.1	0.0	0.0	0.1	0.2	0.1	0.1	0.2	0.3	0.2	0.2	0.4
	50	759	506	522	522	045	871	202	185	509	673	506	587	121	601	342	033
	50	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.0	0.2	0.2	0.2	0.1	0.3
	0	441	294	273	101	659	549	080	746	948	298	972	079	686	238	716	452
	10	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.2	0.1	0.1	0.2
	0	198	132	125	744	348	290	023	400	322	882	537	544	069	724	134	806
	15	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.1	0.0	0.2
	0	140	093	082	604	236	196	011	271	128	752	404	341	842	535	918	525
	20	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.0	0.2
	0	094	063	060	515	178	148	006	207	980	654	332	214	681	401	805	364
2	30	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1	0.2	0.1	0.1	0.2	0.2	0.2	0.1	0.3
	50	762	476	597	527	938	781	160	053	360	475	580	484	485	071	718	454
	50	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.1	0.2	0.1	0.1	0.2
	0	463	290	325	126	603	502	068	687	848	155	013	989	105	755	240	934
	10	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.1	0.0	0.2
	0	238	149	143	747	305	254	018	351	259	787	551	467	578	315	799	355
	15	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.1	0.0	0.2
	0	160	100	095	609	209	174	008	242	038	649	411	267	352	127	650	125
	20	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.0	0.1
	0	125	078	071	526	158	132	005	184	917	573	340	152	241	034	557	967
3	30	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.1	0.1	0.1	0.2
	50	869	511	658	509	864	720	138	978	356	386	656	394	895	579	285	988
	50	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.1	0.1	0.1	0.0	0.2
	0	484	284	350	101	538	449	055	616	761	036	066	921	704	420	995	628
	10	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.1	0.0	0.2
	0	251	148	158	739	278	232	015	321	197	704	570	404	252	043	629	090
	15	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	0	159	094	105	602	191	159	007	222	904	532	406	186	992	827	497	857
	20	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	0	114	067	078	521	143	119	004	167	799	470	326	063	916	763	423	713
4	30	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.1	0.0	0.0	0.2
	50	966	483	931	526	667	556	082	756	200	100	996	234	136	947	654	132
	50	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.0	0.1	0.1	0.0	0.0	0.0	0.1
	0	628	314	508	127	419	349	033	482	613	807	204	735	850	708	467	800
	10	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	0	299	150	230	758	218	182	009	250	986	493	665	290	556	463	305	454
	15	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
	20	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0

50	186	093	147	606	143	119	004	165	732	366	459	071	419	349	241	293
20	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1
00	125	062	107	517	110	092	002	129	590	295	363	952	344	287	208	202

Table 8: Biases and MSEs of $\hat{\alpha}$ and $\hat{\beta}$ using ML and L-M methods at $\beta = 1.5$ and multiple values of α

		Maximum likelihood								L-moments							
		$\hat{\alpha}$				$\hat{\beta}$				$\hat{\alpha}$				$\hat{\beta}$			
α	n	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE
30	30	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.2	0.1	0.2	0.3
	751	501	552	566	857	571	146	804	546	697	576	646	750	834	448	299	
50	50	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.0	0.2	0.2	0.1	0.1	0.2
	447	298	292	140	520	346	054	489	927	285	991	098	204	469	698	747	
100	100	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.1	0.1	0.2
	219	146	132	766	260	173	014	247	386	924	591	621	707	138	155	266	
150	150	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.0	0.2
	167	111	085	615	178	119	006	167	166	777	431	384	461	974	925	028	
200	200	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.1	0.0	0.0	0.1	
	107	071	064	532	130	087	003	123	976	650	347	243	259	839	791	875	
300	300	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.2	0.1	0.1	0.2
	830	519	636	576	788	525	122	738	488	555	696	574	056	370	790	821	
500	500	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.2	0.1	0.1	0.1	0.2
	470	294	337	147	468	312	044	443	829	143	056	031	641	094	430	521	
1000	1000	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.0	0.1
	243	152	153	773	240	160	011	225	245	778	596	525	170	780	880	978	
1500	1500	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.0	0.1
	147	092	096	614	159	106	005	149	024	640	435	303	014	676	714	782	
2000	2000	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1	
	113	070	073	533	119	079	003	112	861	538	365	193	879	586	622	662	
3000	3000	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.1	0.1	0.1	0.2
	827	487	716	574	708	472	102	672	301	354	725	443	536	024	394	489	
5000	5000	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.1	0.1	0.0	0.1	0.2
	506	297	379	145	432	288	037	408	720	012	067	921	207	805	008	116	
10000	10000	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	226	133	171	770	213	142	009	198	138	670	602	443	857	571	662	715	
15000	15000	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1	
	164	097	111	618	144	096	004	135	934	549	435	227	717	478	526	529	
20000	20000	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1	
	129	076	084	539	107	071	002	101	802	472	354	107	616	411	459	429	
30000	30000	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.2	0.2	0.0	0.0	0.0	0.1
	986	493	976	562	541	360	059	510	179	090	098	290	756	504	716	784	

0	5	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.0	0.1	0.1	0.0	0.0	0.0	0.1
	0	593	297	527	148	336	224	023	317	603	801	290	796	612	408	514	512
1	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	0	276	138	246	785	165	110	006	156	947	474	705	328	345	230	352	250
1	5	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	0	182	091	159	630	112	075	003	105	733	366	510	129	267	178	282	119
2	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1
	0	144	072	118	542	083	056	001	079	611	306	395	994	225	150	229	009

Table 9: Biases and MSEs of $\hat{\alpha}$ and $\hat{\beta}$ using ML and L-M methods at $\beta = 2.0$ and multiple values of α

		Maximum likelihood								L-moments							
		$\hat{\alpha}$				$\hat{\beta}$				$\hat{\alpha}$				$\hat{\beta}$			
α	n	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE
3	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.2	0.1	0.3	0.2
	0	767	511	580	605	885	442	160	633	491	660	701	749	687	343	013	744
5	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.2	0.2	0.1	0.2	0.2
	0	464	309	315	182	527	264	057	376	890	260	060	171	182	091	148	317
1	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.1	0.1
	0	211	141	140	789	266	133	014	186	320	880	626	668	663	832	457	909
5	1	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.1	0.1
	0	139	093	091	635	176	088	006	124	104	736	472	449	462	731	180	718
2	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.1	0.0	0.1	0.1
	0	106	071	067	547	133	066	004	095	938	625	392	320	255	628	044	615
3	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.2	0.1	0.2	0.2
	0	846	529	680	629	810	405	133	576	457	536	809	658	075	037	234	363
5	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.2	0.1	0.0	0.1	0.2
	0	512	320	366	195	477	239	046	338	849	156	132	103	679	840	603	002
1	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.1	0.1
	0	250	156	163	797	240	120	012	170	255	785	647	589	230	615	086	648
6	1	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.0	0.1
	0	156	097	107	647	159	080	051	356	002	626	485	376	009	504	862	468
2	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	0	114	071	077	547	119	060	003	084	878	548	393	239	919	459	746	366
3	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.1	0.0	0.1	0.2
	0	931	547	768	630	735	367	108	520	447	440	865	541	661	830	713	069
5	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.2	0.1	0.0	0.1	0.1
	0	518	305	412	194	440	220	039	313	780	047	170	012	321	660	222	748
1	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	0	258	152	186	803	216	108	009	152	175	691	646	495	900	450	814	426
7	1	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	0	178	105	119	642	145	072	004	103	930	547	470	275	695	348	645	270

2	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	134	079	088	553	108	054	002	077	803	472	371	133	639	319	529	150
3	0.1	0.0	0.1	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.2	0.2	0.0	0.0	0.0	0.1
	085	543	100	658	580	290	068	411	268	134	244	368	830	415	915	512
5	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.0	0.1	0.1	0.0	0.0	0.0	0.1
	607	303	574	198	350	175	025	248	582	791	379	856	620	310	658	283
1	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	300	150	258	802	171	086	006	120	997	498	748	367	401	200	413	016
0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0
	228	114	179	669	112	056	003	081	794	397	547	169	316	158	314	885
2	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0
	145	072	124	557	088	044	002	061	624	312	429	035	267	133	259	805

Table 10: Biases and MSEs of $\hat{\alpha}$ and $\hat{\beta}$ using ML and L-M methods at $\beta = 2.5$ and multiple values of α

		Maximum likelihood								L-moments							
		$\hat{\alpha}$				$\hat{\beta}$				$\hat{\alpha}$				$\hat{\beta}$			
α	n	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE
3	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.3	0.1	0.3	0.2
	5	825	550	642	689	983	393	198	562	577	718	776	809	001	200	988	526
5	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.2	0.2	0.0	0.2	0.2
	5	462	308	323	198	578	231	068	329	948	299	103	214	482	993	812	121
1	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.1	0.1
	5	219	146	146	805	292	117	017	165	367	911	641	687	884	754	915	750
5	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.1	0.1
	5	160	107	098	658	194	078	007	109	088	725	480	461	490	596	552	576
2	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.1	0.0	0.1	0.1
	5	117	078	071	562	143	057	004	081	990	660	400	334	417	567	317	452
3	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.2	0.0	0.2	0.2
	5	803	502	703	657	891	356	161	507	409	506	795	648	304	922	893	151
5	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.2	0.1	0.0	0.2	0.1
	5	510	319	380	218	533	213	058	303	865	165	177	144	892	757	488	995
1	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.1	0.1
	5	253	158	171	818	262	105	014	148	262	789	658	603	339	535	398	495
6	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.1	0.0	0.1	0.1
	5	160	100	110	654	176	070	006	100	994	621	492	386	090	436	145	353
2	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1
	5	132	083	083	568	132	053	003	075	860	538	407	261	945	378	973	248
1	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.2	0.2	0.1	0.0	0.2	0.1
	5	974	573	829	693	818	327	135	465	540	494	025	647	876	750	240	893
7	0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.2	0.1	0.0	0.1	0.1
	5	542	319	433	224	492	197	048	278	821	071	200	038	492	597	503	551

1	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.1	0.1	
0	284	167	201	834	240	096	012	138	212	713	685	539	014	406	049	296	
1	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1	
5	173	102	125	659	162	065	005	092	908	534	486	297	774	310	836	156	
2	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.1	
0	130	077	092	564	120	048	003	068	794	467	388	159	706	283	678	042	
3	0.1	0.0	0.1	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.2	0.2	0.0	0.0	0.1	0.1	
0	031	515	125	677	655	262	087	372	207	103	304	400	918	367	204	388	
5	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.0	0.1	0.1	0.0	0.0	0.0	0.1	
0	671	336	645	270	389	156	031	222	635	818	444	900	702	281	805	135	
1	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	
2	0	295	148	278	834	196	078	008	110	972	486	759	378	448	179	513	906
0	1	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	
1	5	212	106	177	665	130	052	003	074	750	375	536	157	342	137	368	767
2	0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	
0	0	157	079	133	577	098	039	002	055	639	319	437	045	309	123	310	704

Table 11: Biases and MSEs of $\hat{\alpha}$ and $\hat{\beta}$ using ML and L-M methods at $\beta = 3.0$, $\beta = 5.0$ and $\alpha = 1.5$, $\alpha = 2.0$

		$\beta = 3.0$															
		Maximum likelihood								L-moments							
		$\hat{\alpha}$				$\hat{\beta}$				$\hat{\alpha}$				$\hat{\beta}$			
α	n	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE	Bias	R.Bias	MSE	SMSE
1.5	30	0.0	0.0	0.0	0.1	0.1	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.3	0.1	0.4	0.2
	50	832	555	629	671	092	364	244	521	608	739	746	786	327	109	817	314
	100	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.2	0.2	0.0	0.3	0.1
	200	486	324	342	232	654	218	087	311	983	322	103	214	747	916	510	975
	300	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.2	0.0	0.2	0.1
	400	243	162	156	833	325	108	021	153	411	940	631	675	111	704	315	604
	500	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.1	0.1
	600	144	096	096	654	215	072	009	102	138	759	466	439	787	596	906	455
	700	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.1	0.0	0.1	0.1
	800	133	089	074	573	164	055	005	077	033	688	389	315	607	536	631	346
	900	0.1	0.0	0.1	0.1	0.0	0.0	0.0	0.0	0.2	0.1	0.2	0.2	0.1	0.0	0.1	0.1
	1000	058	529	171	711	741	247	112	352	272	136	300	398	133	378	417	255
2.0	30	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.0	0.1	0.1	0.0	0.0	0.0	0.1
	50	710	355	663	287	443	148	040	209	751	876	435	894	916	305	935	019
	100	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.0	0.0	0.0	0.0
	200	322	161	291	853	219	073	010	103	093	547	725	346	642	214	571	796
	300	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0
	400	187	094	183	677	144	048	004	067	860	430	512	131	550	183	441	700
	500	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0
	600	154	077	138	588	110	037	002	052	740	370	429	036	465	155	404	670
	700	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0
	800	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0
	900	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0
	1000	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0
		$\beta = 5.0$															
1.5	30	0.0	0.0	0.0	0.1	0.1	0.0	0.0	0.0	0.2	0.1	0.1	0.2	0.4	0.0	1.0	0.2
	50	838	559	686	747	625	325	504	449	679	786	829	851	992	998	850	083
	100	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.1	0.1	0.2	0.3	0.0	0.8	0.1
	200	508	338	360	265	973	195	192	277	997	331	159	269	729	746	598	854
	300	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.2	0.0	0.5	0.1
	400	235	157	170	868	478	096	046	136	398	932	635	680	959	592	290	455
	500	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.2	0.0	0.4	0.1
	600	163	109	106	687	315	063	020	090	137	758	491	477	350	470	828	390
	700	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.2	0.0	0.4	0.1
	800	137	091	078	588	237	047	011	067	007	671	400	333	100	420	091	279
	900	0.1	0.0	0.1	0.1	0.1	0.0	0.0	0.0	0.2	0.1	0.2	0.2	0.1	0.0	0.3	0.1
	1000	237	618	329	823	095	219	241	311	496	248	602	550	647	329	374	162
2.0	30	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0	0.1	0.0	0.1	0.1	0.1	0.0	0.2	0.0
	50	714	357	690	313	671	134	092	192	698	849	538	961	178	236	425	985
	100	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.0	0.0	0.1	0.0
	200	307	153	299	864	326	065	021	091	036	518	806	420	811	162	435	758

1	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.1	0.0
5	210	105	200	708	218	044	010	062	742	371	575	199	543	109	199	692
0																
2	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.1	0.0	0.0	0.0	0.0
0	174	087	152	616	162	032	005	046	641	321	447	058	490	098	865	588
0																

Table 12: Biases and MSEs of $\hat{\alpha}$ and $\hat{\beta}$ using the three estimation methods at $\beta = 1.5, 2.0$ and $\alpha = 2.5, 3.0, 5.0$

$\beta = 1.5$													
		Maximum likelihood				L-moments				Moments			
		$\hat{\alpha}$		$\hat{\beta}$		$\hat{\alpha}$		$\hat{\beta}$		$\hat{\alpha}$		$\hat{\beta}$	
α	n	Bias	MSE	Bias	MSE	Bias	MSE	Bias	MSE	Bias	MSE	Bias	MSE
2.5	30	0.1277	0.1573	0.0396	0.0031	0.2064	0.2879	0.0259	0.0356	0.7650	0.9387	0.4133	0.2326
	50	0.0727	0.0848	0.0241	0.0011	0.1367	0.1787	0.0173	0.0260	0.6109	0.6183	0.3597	0.1797
	100	0.0344	0.0392	0.0118	0.0003	0.0793	0.0906	0.0109	0.0148	0.4722	0.3569	0.3187	0.1288
	150	0.0252	0.0258	0.0081	0.0001	0.0567	0.0648	0.0066	0.0110	0.4069	0.2713	0.2853	0.1048
	200	0.0165	0.0185	0.0060	0.0001	0.0422	0.0498	0.0046	0.0086	0.3690	0.2255	0.2681	0.0935
3.0	30	0.1488	0.2340	0.0306	0.0019	0.2039	0.3939	0.0122	0.0195	0.7312	1.0614	0.2332	0.0850
	50	0.0975	0.1304	0.0182	0.0007	0.1400	0.2404	0.0068	0.0131	0.5693	0.6736	0.1912	0.0614
	100	0.0469	0.0585	0.0091	0.0002	0.0772	0.1210	0.0045	0.0072	0.4030	0.3774	0.1489	0.0428
	150	0.0318	0.0377	0.0060	0.0001	0.0526	0.0820	0.0024	0.0050	0.3324	0.2710	0.1280	0.0349
	200	0.0234	0.0282	0.0045	0.0000	0.0423	0.0656	0.0022	0.0039	0.2926	0.2326	0.1136	0.0324
5.0	30	0.2852	0.7512	0.0149	0.0004	0.2689	1.0441	0.0019	0.0036	0.7831	2.0364	0.0520	0.0110
	50	0.1716	0.4021	0.0090	0.0002	0.1778	0.6150	0.0021	0.0021	0.5533	1.2363	0.0382	0.0082
	100	0.0812	0.1731	0.0045	0.0000	0.0842	0.2831	0.0008	0.0011	0.3261	0.6231	0.0234	0.0057
	150	0.0549	0.1150	0.0030	0.0000	0.0601	0.1908	0.0007	0.0001	0.2488	0.4372	0.0184	0.0043
	200	0.0434	0.0828	0.0023	0.0000	0.0485	0.1422	0.0006	0.0005	0.2034	0.3410	0.0149	0.0037
$\beta = 2.0$													
2.5	30	0.1408	0.1825	0.0424	0.0036	0.2226	0.3221	0.0330	0.0422	0.7986	1.0225	0.4732	0.3032
	50	0.0794	0.0958	0.0257	0.0014	0.1472	0.1918	0.0236	0.0287	0.6366	0.6555	0.4143	0.2325
	100	0.0369	0.0421	0.0128	0.0003	0.0784	0.0963	0.0119	0.0172	0.4754	0.3783	0.3395	0.1650
	150	0.0258	0.0274	0.0085	0.0001	0.0574	0.0683	0.0083	0.0128	0.4129	0.2887	0.3060	0.1377
	200	0.0200	0.0202	0.0062	0.0001	0.0506	0.0521	0.0091	0.0092	0.3783	0.2416	0.2859	0.1219
3.0	30	0.1763	0.2785	0.0329	0.0022	0.2362	0.4536	0.0171	0.0218	0.7828	1.2046	0.2590	0.1148
	50	0.0990	0.1458	0.0193	0.0008	0.1468	0.2577	0.0112	0.0147	0.5893	0.7210	0.2156	0.0868
	100	0.0477	0.0634	0.0098	0.0002	0.0812	0.1294	0.0069	0.0083	0.4181	0.4012	0.1655	0.0622
	150	0.0317	0.0407	0.0065	0.0001	0.0582	0.0874	0.0055	0.0055	0.3410	0.2976	0.1393	0.0531
	200	0.0244	0.0308	0.0049	0.0001	0.0453	0.0682	0.0041	0.0045	0.2992	0.2390	0.1264	0.0449
5.0	30	0.2799	0.8257	0.0173	0.0006	0.2636	1.1457	0.0027	0.0049	0.8171	2.3183	0.0647	0.0146
	50	0.1691	0.4313	0.0103	0.0002	0.1683	0.6419	0.0020	0.0028	0.5673	1.3297	0.0463	0.0107
	100	0.0835	0.2037	0.0052	0.0001	0.0882	0.3272	0.0012	0.0015	0.3388	0.7019	0.0282	0.0072
	150	0.0524	0.1257	0.0035	0.0000	0.0552	0.2031	0.0007	0.0009	0.2512	0.4690	0.0220	0.0055
	200	0.0431	0.0952	0.0026	0.0000	0.0507	0.1567	0.0011	0.0007	0.2095	0.3759	0.0177	0.0050

Table 13: Biases and MSEs of $\hat{\alpha}$ and $\hat{\beta}$ using the three estimation methods at $\beta = 2.5, 3.0$ and $\alpha = 2.5, 3.0, 5.0$

$\beta = 2.5$													
		Maximum likelihood				L-moments				Moments			
		$\hat{\alpha}$		$\hat{\beta}$		$\hat{\alpha}$		$\hat{\beta}$		$\hat{\alpha}$		$\hat{\beta}$	
α	n	Bias	MSE	Bias	MSE	Bias	MSE	Bias	MSE	Bias	MSE	Bias	MSE
2.5	30	0.1429	0.1949	0.0488	0.0048	0.2272	0.3275	0.0419	0.0536	0.8137	1.0359	0.5586	0.4134
	50	0.0867	0.1018	0.0291	0.0017	0.1544	0.1979	0.0283	0.0364	0.6479	0.6653	0.4813	0.2995
	100	0.0379	0.0460	0.0143	0.0004	0.0821	0.1000	0.0161	0.0209	0.4864	0.3817	0.4044	0.2098
	150	0.0271	0.0296	0.0095	0.0002	0.0608	0.0721	0.0107	0.0155	0.4198	0.2911	0.3606	0.1709
	200	0.0196	0.0217	0.0072	0.0001	0.0483	0.0552	0.0092	0.0119	0.3826	0.2434	0.3374	0.1521
3.0	30	0.1726	0.2770	0.0385	0.0030	0.2321	0.4553	0.0206	0.0294	0.7888	1.2008	0.3186	0.1464
	50	0.1039	0.1512	0.0226	0.0010	0.1519	0.2692	0.0133	0.0194	0.6050	0.7469	0.2575	0.1068
	100	0.0490	0.0690	0.0114	0.0003	0.0824	0.1362	0.0083	0.0108	0.4239	0.4052	0.1969	0.0744
	150	0.0322	0.0444	0.0076	0.0001	0.0534	0.0930	0.0042	0.0076	0.3439	0.2920	0.1662	0.0618
	200	0.0259	0.0322	0.0057	0.0006	0.0450	0.0706	0.0038	0.0056	0.3058	0.2382	0.1492	0.0548
5.0	30	0.3187	0.9226	0.0207	0.0009	0.3110	1.2703	0.0045	0.0067	0.8810	2.5868	0.0770	0.0223
	50	0.1763	0.4697	0.0125	0.0003	0.1765	0.6932	0.0028	0.0038	0.5828	1.4232	0.0553	0.0150
	100	0.0885	0.2040	0.0062	0.0008	0.0930	0.3248	0.0016	0.0020	0.3506	0.0347	0.7033	0.0095
	150	0.0670	0.1400	0.0040	0.0000	0.0722	0.2282	0.0009	0.0013	0.2741	0.5207	0.0260	0.0079
	200	0.0404	0.0994	0.0030	0.0000	0.0433	0.1621	0.0007	0.0010	0.2057	0.3838	0.0209	0.0067
$\beta = 3.0$													
2.5	30	0.1515	0.1982	0.0556	0.0064	0.2334	0.3433	0.0446	0.0691	0.8185	1.0708	0.6231	0.5238
	50	0.0786	0.1025	0.0332	0.0023	0.1482	0.1994	0.0330	0.0462	0.6480	0.6759	0.5514	0.3957
	100	0.0411	0.0474	0.0164	0.0005	0.0868	0.1056	0.0182	0.0270	0.4933	0.3955	0.4574	0.2718
	150	0.0310	0.0309	0.0109	0.0002	0.0634	0.0734	0.0116	0.0197	0.4233	0.2984	0.4048	0.2212
	200	0.0204	0.0226	0.0082	0.0001	0.0493	0.0577	0.0105	0.0163	0.3847	0.2498	0.3799	0.1996
3.0	30	0.1798	0.2964	0.0431	0.0038	0.2462	0.4729	0.0272	0.0379	0.8073	1.2550	0.3548	0.1964
	50	0.1034	0.1539	0.0264	0.0014	0.1546	0.2715	0.0179	0.0246	0.6101	0.7577	0.2952	0.1429
	100	0.0558	0.0720	0.0132	0.0004	0.0874	0.1433	0.0086	0.0144	0.4304	0.4213	0.2216	0.0967
	150	0.0327	0.0457	0.0087	0.0002	0.0574	0.0960	0.0063	0.0096	0.3508	0.3032	0.1905	0.0831
	200	0.0263	0.0337	0.0066	0.0001	0.0472	0.0717	0.0057	0.0072	0.3081	0.2427	0.1699	0.0724
5.0	30	0.3036	0.9192	0.0237	0.0011	0.2990	1.2940	0.0050	0.0091	0.8409	2.5206	0.0838	0.0331
	50	0.1812	0.4716	0.0144	0.0004	0.1830	0.7025	0.0032	0.0053	0.5943	1.4610	0.0645	0.0202
	100	0.0898	0.2190	0.0071	0.0001	0.0955	0.3458	0.0017	0.0027	0.3583	0.7400	0.0403	0.0140
	150	0.0536	0.1371	0.0048	0.0001	0.0593	0.2253	0.0012	0.0017	0.2635	0.5071	0.0312	0.0104
	200	0.0441	0.1030	0.0035	0.0000	0.0469	0.1694	0.0008	0.0014	0.1990	0.3856	0.0218	0.0097

Table 14: R.Biases and SMSEs of $\hat{\alpha}$ and $\hat{\beta}$ using the three estimation methods at $\beta = 1.5, 2.0$ and $\alpha = 2.5, 3.0, 5.0$

$\beta = 1.5$													
		Maximum likelihood				L-moments				Moments			
		$\hat{\alpha}$		$\hat{\beta}$		$\hat{\alpha}$		$\hat{\beta}$		$\hat{\alpha}$		$\hat{\beta}$	
α	n	R.Bias	SMSE	R.Bias	SMSE	R.Bias	SMSE	R.Bias	SMSE	R.Bias	SMSE	R.Bias	SMSE
2.5	30	0.0511	0.1587	0.0264	0.0371	0.0825	0.2146	0.0173	0.1258	0.3060	0.3876	0.2755	0.3215
	50	0.0291	0.1165	0.0161	0.0224	0.0547	0.1691	0.0115	0.1076	0.2444	0.3145	0.2398	0.2826
	100	0.0138	0.0792	0.0079	0.0112	0.0317	0.1204	0.0073	0.0810	0.1889	0.2389	0.2125	0.2393
	150	0.0101	0.0643	0.0054	0.0076	0.0227	0.1018	0.0044	0.0700	0.1628	0.2083	0.1902	0.2158
	200	0.0066	0.0544	0.0040	0.0056	0.0169	0.0892	0.0030	0.0618	0.1476	0.1900	0.1787	0.2038
3.0	30	0.0496	0.1612	0.0204	0.0289	0.0680	0.2092	0.0081	0.0932	0.2437	0.3434	0.1555	0.1943
	50	0.0325	0.1204	0.0121	0.0171	0.0467	0.1634	0.0045	0.0762	0.1898	0.2736	0.1275	0.1652
	100	0.0156	0.0806	0.0060	0.0084	0.0257	0.1160	0.0030	0.0564	0.1343	0.2048	0.0993	0.1379
	150	0.0106	0.0647	0.0040	0.0056	0.0175	0.0954	0.0016	0.0471	0.1108	0.1735	0.0853	0.1245
	200	0.0078	0.0559	0.0030	0.0042	0.0141	0.0854	0.0015	0.0418	0.0975	0.1608	0.0757	0.1200
5.0	30	0.0570	0.1733	0.0099	0.0140	0.0538	0.2044	0.0013	0.0399	0.1566	0.2854	0.0347	0.0698
	50	0.0343	0.1268	0.0060	0.0084	0.0356	0.1568	0.0014	0.0308	0.1107	0.2224	0.0254	0.0603
	100	0.0162	0.0832	0.0030	0.0042	0.0168	0.1064	0.0005	0.0219	0.0652	0.1579	0.0156	0.0503
	150	0.0110	0.0678	0.0020	0.0030	0.0120	0.0874	0.0005	0.0056	0.0498	0.1322	0.0123	0.0439
	200	0.0087	0.0576	0.0015	0.0021	0.0097	0.0754	0.0004	0.0153	0.0407	0.1168	0.0099	0.0404
$\beta = 2.0$													
2.5	30	0.0563	0.1709	0.0212	0.0301	0.0891	0.2270	0.0165	0.1028	0.3194	0.4045	0.2366	0.2753
	50	0.0317	0.1238	0.0128	0.0184	0.0589	0.1752	0.0118	0.0847	0.2546	0.3238	0.2072	0.2411
	100	0.0148	0.0820	0.0064	0.0089	0.0314	0.1241	0.0059	0.0655	0.1901	0.2460	0.1698	0.2031
	150	0.0103	0.0662	0.0043	0.0059	0.0230	0.1046	0.0041	0.0565	0.1652	0.2149	0.1530	0.1855
	200	0.0080	0.0569	0.0031	0.0045	0.0202	0.0913	0.0046	0.0479	0.1513	0.1966	0.1430	0.1746
3.0	30	0.0588	0.1759	0.0165	0.0233	0.0787	0.2245	0.0085	0.0738	0.2609	0.3658	0.1295	0.1694
	50	0.0330	0.1273	0.0097	0.0138	0.0489	0.1692	0.0056	0.0607	0.1964	0.2830	0.1078	0.1473
	100	0.0159	0.0839	0.0049	0.0069	0.0271	0.1199	0.0034	0.0456	0.1394	0.2111	0.0828	0.1247
	150	0.0106	0.0672	0.0033	0.0047	0.0194	0.0985	0.0028	0.0371	0.1137	0.1818	0.0696	0.1152
	200	0.0081	0.0585	0.0025	0.0035	0.0151	0.0870	0.0020	0.0334	0.0997	0.1630	0.0632	0.1060
5.0	30	0.0560	0.1817	0.0087	0.0123	0.0527	0.2141	0.0014	0.0348	0.1634	0.3045	0.0323	0.0603
	50	0.0338	0.1313	0.0052	0.0072	0.0337	0.1602	0.0010	0.0264	0.1135	0.2306	0.0232	0.0516
	100	0.0167	0.0903	0.0026	0.0039	0.0176	0.1144	0.0006	0.0191	0.0678	0.1676	0.0141	0.0425
	150	0.0105	0.0709	0.0017	0.0022	0.0110	0.0901	0.0004	0.0153	0.0502	0.1370	0.0110	0.0371
	200	0.0086	0.0617	0.0013	0.0016	0.0101	0.0792	0.0005	0.0133	0.0419	0.1226	0.0088	0.0354

Table 15: R.Biases and SMSEs of $\hat{\alpha}$ and $\hat{\beta}$ using the three estimation methods at $\beta = 2.5, 3.0$ and $\alpha = 2.5, 3.0, 5.0$

$\beta = 2.5$													
		Maximum likelihood				L-moments				Moments			
		$\hat{\alpha}$		$\hat{\beta}$		$\hat{\alpha}$		$\hat{\beta}$		$\hat{\alpha}$		$\hat{\beta}$	
α	n	R.Bias	SMSE	R.Bias	SMSE	R.Bias	SMSE	R.Bias	SMSE	R.Bias	SMSE	R.Bias	SMSE
2.5	30	0.0572	0.1766	0.0195	0.0277	0.0909	0.2289	0.0167	0.0926	0.3255	0.4071	0.2234	0.2572
	50	0.0347	0.1276	0.0116	0.0165	0.0618	0.1780	0.0113	0.0763	0.2592	0.3263	0.1925	0.2189
	100	0.0152	0.0858	0.0057	0.0081	0.0328	0.1265	0.0064	0.0578	0.1946	0.2471	0.1617	0.1832
	150	0.0108	0.0688	0.0038	0.0054	0.0243	0.1074	0.0043	0.0498	0.1679	0.2158	0.1442	0.1654
	200	0.0078	0.0589	0.0029	0.0040	0.0193	0.0940	0.0037	0.0437	0.1530	0.1974	0.1350	0.1560
3.0	30	0.0575	0.1754	0.0154	0.0219	0.0774	0.2249	0.0082	0.0686	0.2629	0.3653	0.1274	0.1531
	50	0.0346	0.1296	0.0090	0.0128	0.0506	0.1729	0.0053	0.0558	0.2017	0.2881	0.1030	0.1307
	100	0.0163	0.0875	0.0046	0.0064	0.0275	0.1230	0.0033	0.0415	0.1413	0.2122	0.0788	0.1091
	150	0.0107	0.0702	0.0030	0.0042	0.0178	0.1017	0.0017	0.0348	0.1146	0.1801	0.0665	0.0994
	200	0.0086	0.0598	0.0023	0.0100	0.0150	0.0886	0.0015	0.0300	0.1019	0.1627	0.0597	0.0936
5.0	30	0.0637	0.1921	0.0083	0.0119	0.0622	0.2254	0.0018	0.0328	0.1762	0.3217	0.0308	0.0597
	50	0.0353	0.1371	0.0050	0.0072	0.0353	0.1665	0.0011	0.0248	0.1166	0.2386	0.0221	0.0489
	100	0.0177	0.0903	0.0025	0.0113	0.0186	0.1140	0.0006	0.0177	0.0701	0.0372	0.2813	0.0390
	150	0.0134	0.0748	0.0016	0.0022	0.0144	0.0955	0.0004	0.0145	0.0548	0.1443	0.0104	0.0356
	200	0.0081	0.0631	0.0012	0.0018	0.0087	0.0805	0.0003	0.0125	0.0411	0.1239	0.0083	0.0327
$\beta = 3.0$													
2.5	30	0.0606	0.1781	0.0185	0.0266	0.0934	0.2344	0.0149	0.0876	0.3274	0.4139	0.2077	0.2413
	50	0.0315	0.1281	0.0111	0.0158	0.0593	0.1786	0.0110	0.0717	0.2592	0.3289	0.1838	0.2097
	100	0.0164	0.0871	0.0055	0.0077	0.0347	0.1300	0.0061	0.0548	0.1973	0.2516	0.1525	0.1738
	150	0.0124	0.0703	0.0036	0.0052	0.0254	0.1084	0.0039	0.0468	0.1693	0.2185	0.1349	0.1568
	200	0.0081	0.0601	0.0027	0.0038	0.0197	0.0961	0.0035	0.0426	0.1539	0.1999	0.1266	0.1489
3.0	30	0.0599	0.1815	0.0144	0.0204	0.0821	0.2292	0.0091	0.0649	0.2691	0.3734	0.1183	0.1477
	50	0.0345	0.1308	0.0088	0.0125	0.0515	0.1737	0.0060	0.0522	0.2034	0.2902	0.0984	0.1260
	100	0.0186	0.0894	0.0044	0.0062	0.0291	0.1262	0.0029	0.0400	0.1435	0.2164	0.0739	0.1037
	150	0.0109	0.0713	0.0029	0.0041	0.0191	0.1033	0.0021	0.0327	0.1169	0.1835	0.0635	0.0961
	200	0.0088	0.0612	0.0022	0.0032	0.0157	0.0892	0.0019	0.0282	0.1027	0.1642	0.0566	0.0897
5.0	30	0.0607	0.1917	0.0079	0.0112	0.0598	0.2275	0.0017	0.0317	0.1682	0.3175	0.0279	0.0606
	50	0.0362	0.1373	0.0048	0.0067	0.0366	0.1676	0.0011	0.0243	0.1189	0.2417	0.0215	0.0473
	100	0.0180	0.0936	0.0024	0.0033	0.0191	0.1176	0.0006	0.0172	0.0717	0.1720	0.0134	0.0394
	150	0.0107	0.0741	0.0016	0.0024	0.0119	0.0949	0.0004	0.0139	0.0527	0.1424	0.0104	0.0340
	200	0.0088	0.0642	0.0012	0.0015	0.0094	0.0823	0.0003	0.0124	0.0398	0.1242	0.0073	0.0329